

Uni-variate Statistics

- Types of Random Variables
- Description of RV's
 - using Probability Distributions
 - using Parameters
- Selected Probability Distributions

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Types of Quantitative RV's

- Discrete RV
 - assume values that do not form a continuous range, e.g., number of attendants
 - Its possible values are not limited to only integers.
- Continuous RV's
 - assume a continue range of values, e.g., temperature

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Values of Random Variables

- Quantitative, e.g., income, rainfalls
- Qualitative, e.g., color
- Semi-quantitative or semi-qualitative, e.g.,
course letter grade, income group
 - express order not magnitude

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Probability Distributions

- Tools to describe the Uncertainty of a quantitative random variable
- Popular tools
 - Probability Mass Function (PMF)
 - Probability Density Function (pdf)
 - Cumulative Distribution Function (CDF)

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Estimation and Statistical Inference

Why?

- Unknown real probability distribution
- Unknown parameters
- Estimation accuracy

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Graphing data

- to show the distribution
- type of graphs, e.g.,
 - histogram
 - line graphs
 - bar graphs
 - pie chart

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Data Sampling

- Sampling/observing
 - sampling schemes
 - sample size or number of observations
- Presentation of a data sample
 - graph
 - Calculated statistics

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Calculated Statistics

- to estimate the unknown parameters
- Popular statistics
 - sample mean
 - sample variance
 - sample skewedness
 - sample kurtosis(peakedness or tail-thickness)

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Random Sampling (1)

- Realization of the previous observation will not influence the results of the next observations
- Sample of size n

$$\{X_1, X_2, \dots, X_n\}$$

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Sample Mean

- Estimator of population mean (μ_X)

$$\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$$

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Random Sampling (2)

- X_i 's are random variables having identical probability distributions.
- X_i 's are independent from each other.
- Values of X_i does not influence Values of X_j for $i \neq j$

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Sample Variance

- Estimator of population variance (σ_X^2)

$$s_X^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n - 1}$$

- Why $n-1$ not n ?

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Sample Standard Deviation

- Estimator of population SD or σ_X

$$s_X = \sqrt{s_X^2}$$

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Sample Kurtosis (Peakedness)

- Estimator of population peakedness

$$K = \left\{ \frac{n(n+1)}{(n-1)(n-2)(n-3)} \sum_{i=1}^n \left(\frac{X_i - \bar{X}}{s} \right)^4 \right\} - \frac{3(n-1)^2}{(n-2)(n-3)}$$

- Why?

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Sample Skewedness

- Estimator of population skewedness

$$s_k = \frac{n}{(n-1)(n-2)} \sum_{i=1}^n \left(\frac{X_i - \bar{X}}{s_X} \right)^3$$

- Why?

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Sampling Distributions

- Probability of Calculated Statistics
- Note that they are random
- Assumptions
 - 1) $X \sim N(\mu, \sigma^2)$
 $E(X) = \mu$
 $V(X) = \sigma^2$
 - 2) Random sampling

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Distribution of \bar{X}

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$E(\bar{X}) = \mu$$

$$V(\bar{X}) = \sigma^2 / n$$

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Unbiased Estimators (1)

- Definition: $\hat{\theta}$ is an unbiased estimator of θ
if $E(\hat{\theta}) = \theta$
- Note that \bar{X} is an unbiased estimator of μ .
That is, $E(\bar{X}) = \mu$

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Distribution of s_X^2

$s_X^2 \sim$ non - standard

$$(n-1) \frac{s_X^2}{\sigma^2} \sim \chi^2(n-1)$$

$$E(s_X^2) = \sigma^2$$

$$V(s_X^2) = \frac{2\sigma^4}{n-1}$$

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Unbiased Estimators (2)

- s_X^2 is an unbiased estimator of σ^2
- Reason for dividing by $(n-1)$ is to make the estimators unbiased.
- Same reason for the formulae of sample skewedness and peakedness

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Desired Properties of an Estimator

- Unbiasedness
- Efficiency or smallest possible variance
- If cannot find such an estimator, the next best is a **consistent** estimator
 - asymptotically unbiased, i.e., bias $\rightarrow 0$ as $n \rightarrow \infty$
 - asymptotically diminishing variance, i.e., variance $\rightarrow 0$ as $n \rightarrow \infty$

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Standardized \bar{X} (1)

- Subtract with its mean and divide by its standard deviation

$$Z_{Cal} = \frac{\bar{X} - \mu}{\sigma / \sqrt{n}} \sim N(0,1)$$

- Note that if σ is known, Z can be calculated given μ

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Estimator and Estimate

- Estimator is a formula to estimate a parameter
- Estimate is a calculated or observed value to be used as a proxy for the parameter

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Standardized \bar{X} (2)

- Subtract with its mean and divide by its sample standard deviation

$$t_{Cal} = \frac{\bar{X} - \mu}{s_X / \sqrt{n}} \sim t(n-1)$$

- Does not need to know σ to calculate t .
- Why?

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Standardized \bar{X} (3)

$$t_{cal} = \frac{\frac{\bar{X} - \mu}{\sigma / \sqrt{n}}}{\sqrt{\frac{(n-1) \frac{S_X^2}{\sigma^2}}{n-1}}} \sim t(n-1)$$

Note: see the creation of t-distributed variable

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Central Limit Theorem for \bar{X} (2)

Z_{cal} is approx. distributed as
a standard normal

t_{cal} is approx. t-distributed.

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Central Limit Theorem for \bar{X} (1)

- Violation of normal distribution of X

but still $E(X) = \mu$

$$V(X) = \sigma^2$$

- \bar{X} is approximately normal when $n \rightarrow \infty$
or $\sqrt{n}(\bar{X} - \mu)$ is asymptotically normal

$$\sqrt{n}(\bar{X} - \mu) \overset{A}{\sim} N(0, \sigma^2)$$

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Confidence Interval for μ

Interval estimator for μ (cf. Point estimator)

Concept

Confidence Interval for $\mu = \bar{X} \pm h$

where h = half-width of the confidence

interval

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Factors Affecting h

Sample size (n). h decreases as n increases.

Confidence Level. h increases as confidence level increases

Note that

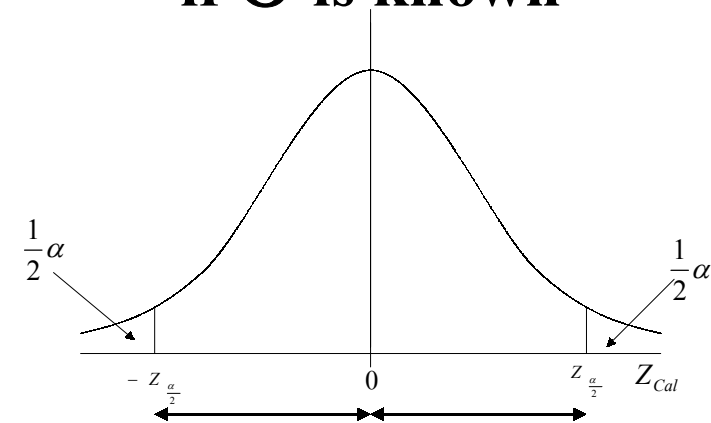
Confidence Level = $(1 - \alpha)100\%$ where α is the significance level (chosen by the one who estimates)

Standard confidence levels are 90%, 95% and 99%

Standard values of α are 0.10, 0.05 and 0.01

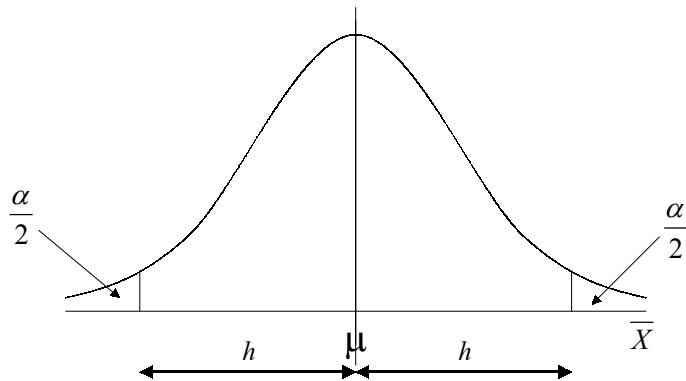
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How to choose h (2) if σ is known



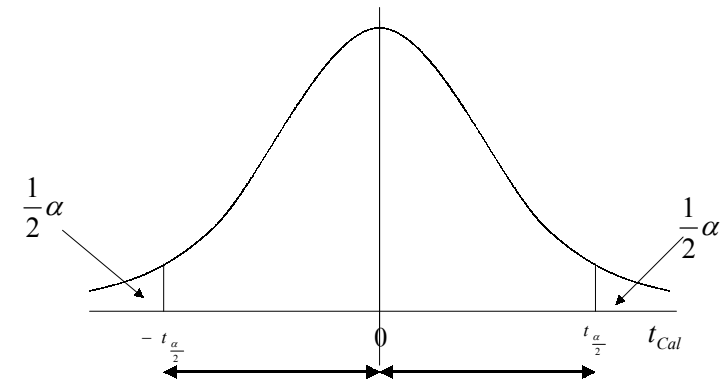
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How to choose h (1)



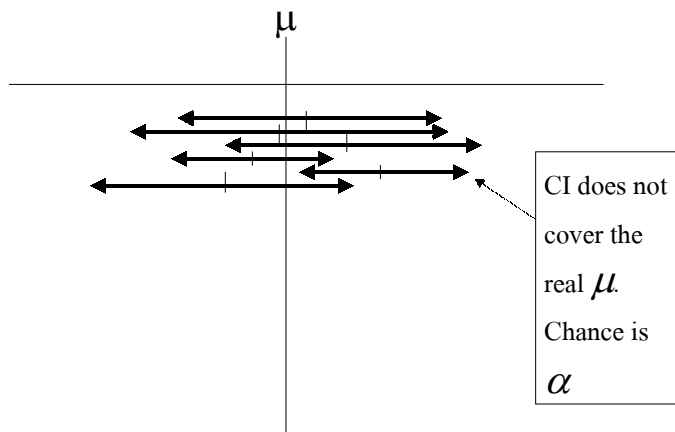
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How to choose h (3) if σ is unknown



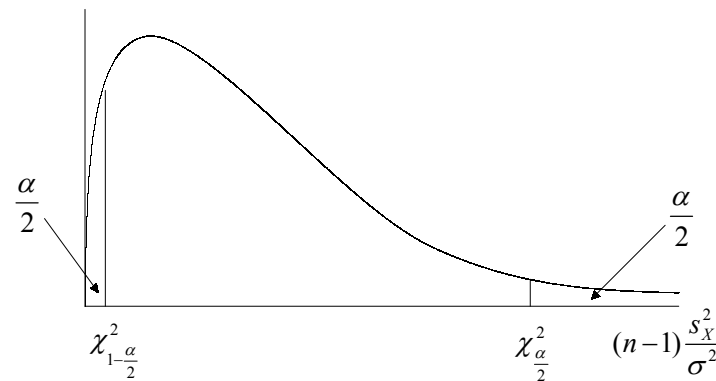
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Interpretation of CI



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How to choose L, U (1)



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Confidence Interval for σ^2

Interval estimator for σ^2 (cf. Point estimator)

Define

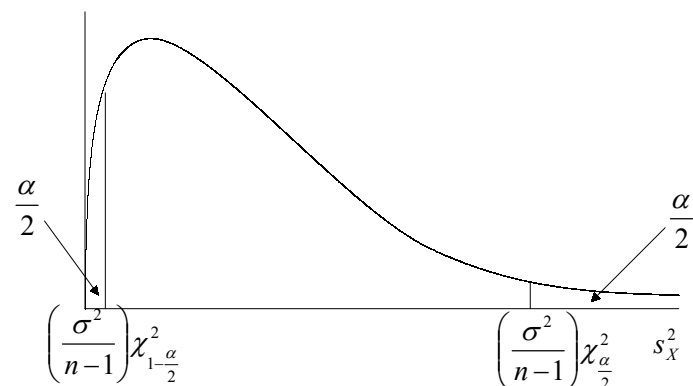
Confidence Interval for $\sigma^2 = \{L, U\}$

where L = lower limit of the confidence interval

U = upper limit of the confidence interval

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How to choose L, U (2)



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How to choose L, U (3)

Note that

$$\left(\frac{\sigma^2}{n-1}\right)\chi^2_{1-\frac{\alpha}{2}} < s_X^2 < \left(\frac{\sigma^2}{n-1}\right)\chi^2_{\frac{\alpha}{2}}$$

Take inverse

$$\left(\frac{n-1}{\sigma^2}\right)\frac{1}{\chi^2_{1-\frac{\alpha}{2}}} > \frac{1}{s_X^2} > \left(\frac{n-1}{\sigma^2}\right)\frac{1}{\chi^2_{\frac{\alpha}{2}}}$$

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How to choose L, U (5)

$(1-\alpha)100\%$ Confidence Interval for

$$\sigma^2 = \left\{ (n-1)\frac{s_X^2}{\chi^2_{\frac{\alpha}{2}}}, (n-1)\frac{s_X^2}{\chi^2_{1-\frac{\alpha}{2}}} \right\}$$

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How to choose L, U (4)

Multiply with $\sigma^2 s_X^2$

$$(n-1)\frac{s_X^2}{\chi^2_{1-\frac{\alpha}{2}}} > \sigma^2 > (n-1)\frac{s_X^2}{\chi^2_{\frac{\alpha}{2}}}$$

Flip over

$$(n-1)\frac{s_X^2}{\chi^2_{\frac{\alpha}{2}}} < \sigma^2 < (n-1)\frac{s_X^2}{\chi^2_{1-\frac{\alpha}{2}}}$$

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One-sided Confidence Interval for σ^2

Concept

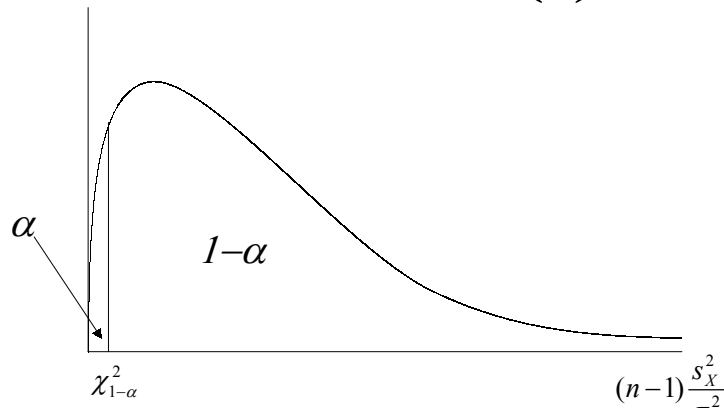
Note that $\sigma^2 \geq 0$ always.

In general, we are more concerned about the upper side of the limit.

One-sided Confidence Interval for $\sigma^2 = \{0, U\}$

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How to choose U for One-sided CI (1)



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How to choose U for One-sided CI (3)

$(1-\alpha)100\%$ One-sided Confidence Interval for

$$\sigma^2 = \left\{ 0, (n-1) \frac{s_X^2}{\chi_{1-\alpha}^2} \right\}$$

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How to choose U for One-sided CI (2)

For $(1-\alpha)100\%$ chance that

$$\chi_{1-\alpha}^2 < (n-1) \frac{s_X^2}{\sigma^2}$$

$$0 < \sigma^2 < (n-1) \frac{s_X^2}{\chi_{1-\alpha}^2}$$

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The Other One-sided Confidence Interval for σ^2

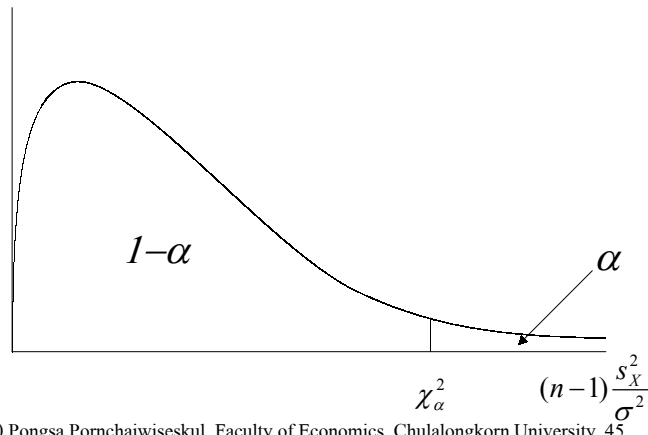
$$= \{L, \infty\}$$

$$= \left\{ (n-1) \frac{s_X^2}{\chi_\alpha^2}, \infty \right\}$$

If we are worried about the lower limit of the confidence interval. How to choose L ?

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How to choose L for One-sided CI



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Two-sided Hypo. Testing for μ (1)

- \bar{X} is an estimator(RV) of μ
- Its observed or calculated value (estimate) could be different from the hypothesized value of μ even though H_0 is true.
- If its observed value is not too far from the hypothesized value, then, we can accept that H_0 is true.

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Hypothesis Testing for μ

Prior info. suggests the null hypothesis H_0 or a hypothesized value of μ

$$H_0 : \mu = 10$$

$$H_1 : \mu \neq 10$$

H_0 is true until a strong enough evidence (observed data) suggests that H_0 is unlikely or, alternatively, H_1 is true.

H_1 is also referred to as the Alternative hypothesis

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Two-sided Hypo. Testing for μ (2)

How far from the hypothesized value is

\bar{X} before H_0 will be rejected?

Criterion

Set the probability of Type I Error equal to the Significant Level (α)

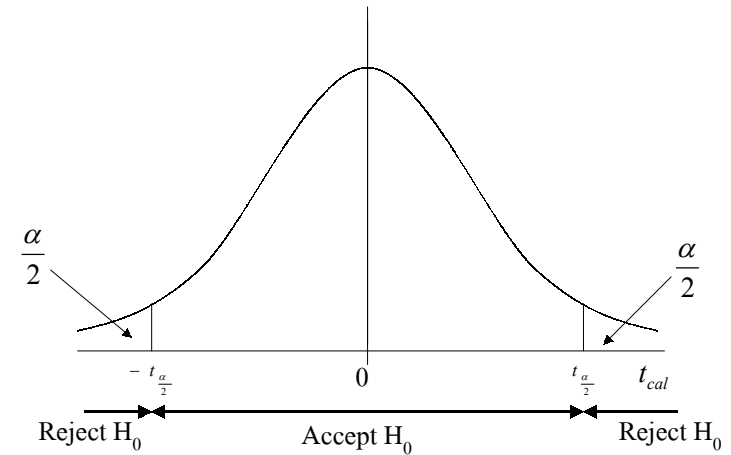
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Type I Error

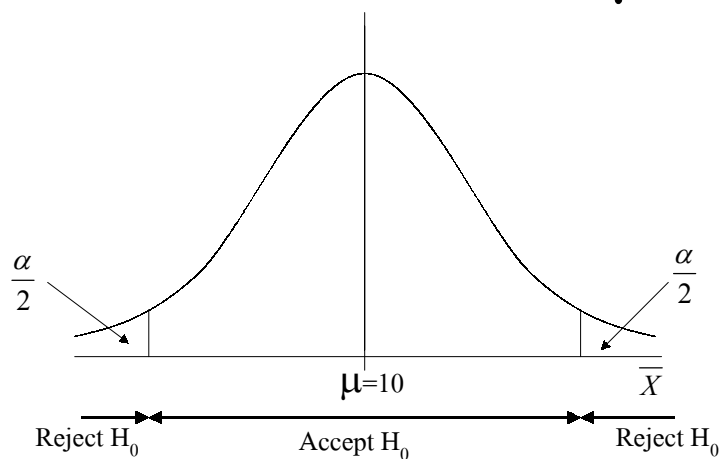
- Type I Error is the event that H_0 has been rejected while it is, in fact, true.
- Low α suggests that the evidence must be very strong before we decide to reject H_0 .
- Remember?
 - significant level is chosen by the tester.
 - standard level of significance

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Two-sided Hypo. Testing for μ (4)



Two-sided Hypo. Testing for μ (3)



Two-sided Hypo. Testing for μ (5)

Calculate $t_{cal} = \frac{\bar{X} - 10}{s_x / \sqrt{n}}$

Criterion compare t_{cal} with t-value from t-table

$$|t_{cal}| \leq t_{\frac{\alpha}{2}} \implies \text{accept } H_0$$

$$|t_{cal}| > t_{\frac{\alpha}{2}} \implies \text{reject } H_0$$

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One-sided Hypo. Testing for μ (1)

Also called One-tailed

Prior info. suggests that no need to be worried about the other tail(say, $\mu < 10$).

Example

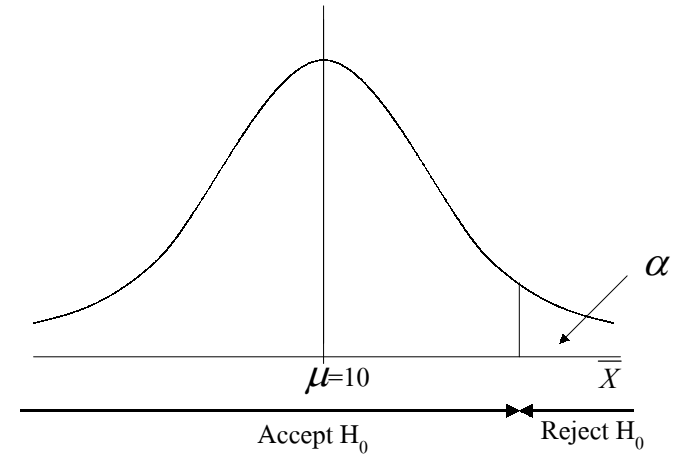
$$H_0: \mu = 10$$

$$H_1: \mu > 10$$

Do a right-tailed test

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One-sided Hypo. Testing for μ (3)



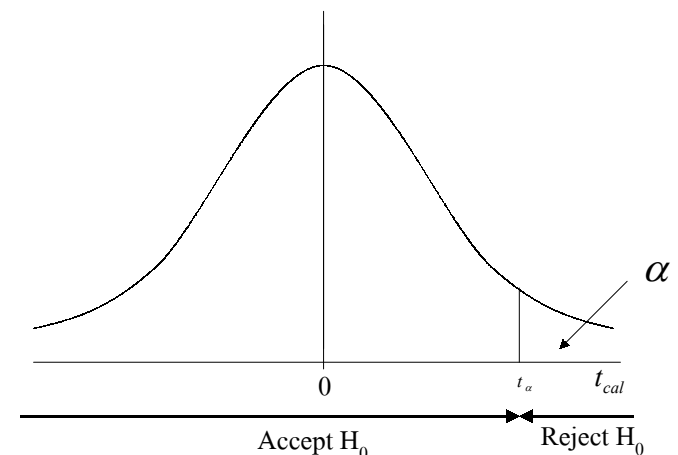
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One-sided Hypo. Testing for μ (2)

- Even if the observed/calculated value is less than or equal to the hypothesized value of μ , there is no good reason to choose $\mu > 10$ over $\mu = 10$
- Only if the observed value is larger than the hypothesized value by a large enough margin, then, we can reject H_0 .

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One-sided Hypo. Testing for μ (4)



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One-sided Hypo. Testing for μ (5)

Calculate $t_{cal} = \frac{\bar{X} - 10}{s_x / \sqrt{n}}$

Criterion compare t_{cal} with t-value

from t-table

$$t_{cal} \leq t_{\alpha} \implies \text{accept } H_0$$

$$t_{cal} > t_{\alpha} \implies \text{reject } H_0$$

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Left-sided Hypo. Testing for μ (2)

Calculate $t_{cal} = \frac{\bar{X} - 10}{s_x / \sqrt{n}}$

Criterion compare t_{cal} with t-value

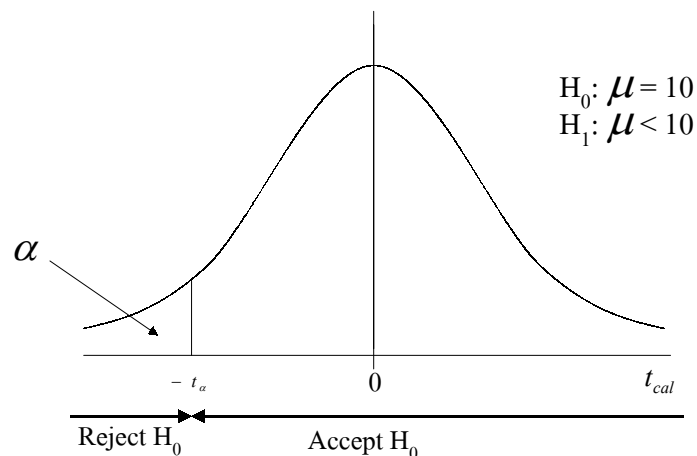
from t-table

$$t_{cal} \geq -t_{\alpha} \implies \text{accept } H_0$$

$$t_{cal} < -t_{\alpha} \implies \text{reject } H_0$$

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Left-sided Hypo. Testing for μ (1)



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Hypothesis Testing for σ^2 (1)

Variance is always non-negative.

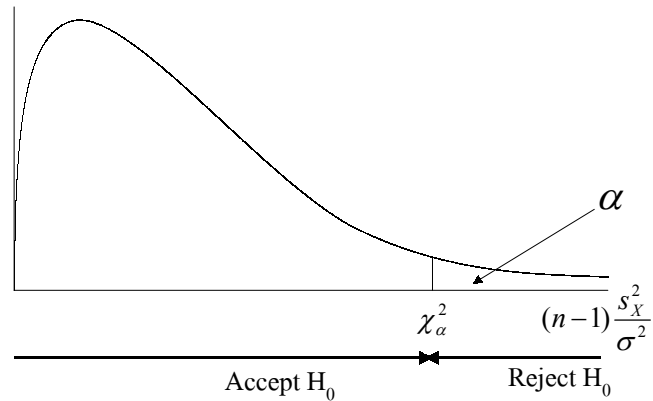
Right-sided test is most interesting since we are more concerned about too high variance, i.e,

$$H_0: \sigma^2 = 16$$

$$H_1: \sigma^2 > 16$$

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Hypothesis Testing for σ^2 (2)



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Hypothesis Testing for σ^2 (3)

Calculate $\chi_{cal}^2 = (n-1) \frac{s_x^2}{16}$

Criterion compare χ_{cal}^2 with χ^2 -value
from χ^2 -table

$$\chi_{cal}^2 \leq \chi_\alpha^2 \implies \text{accept } H_0$$

$$\chi_{cal}^2 > \chi_\alpha^2 \implies \text{reject } H_0$$

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