

Stationary Processes

Types of Stationarity

- Weak
- Strong
- Stochastic

Covered Topics

- Definition of Stationarity
- Types of Time Series
Generating Processes
- Estimation
 - Box-Jenkins Method

Weak Stationarity

$Y \sim$ Weakly Stationary if

- Mean stationary $E(Y_t) = \mu < \infty$ for all t
 \implies time-invariant mean
- Variance stationary $V(Y_t) = \gamma_0 < \infty$ for all t
 \implies time-invariant variance
- Covariance stationary
$$\text{Cov}(Y_t, Y_{t-s}) = \gamma_s < \infty \text{ for all } t, s$$

 \implies time-invariant but spread-specific

Strong Stationarity

Strongly Stationary Process

- require only Covariance Stationarity

Why is Stationarity important?

OLS involving a time series data requires additional assumptions. One of them is “stationarity” of the error terms.

Otherwise, OLS will be invalid.

Stochastic Stationarity

$Y \sim$ Stochastically Stationary if

$$\text{pdf}(\dots, Y_t = a, Y_{t+1} = b, \dots)$$

$$= \text{pdf}(\dots, Y_s = a, Y_{s+1} = b, \dots) \text{ for all } t, s$$

Note that

Weak \Rightarrow Strong
Stochastic \Rightarrow Weak

Stationary MA(q)

$$V(Y_t) = \sigma^2(1 + \theta_1^2 + \theta_2^2 + \dots + \theta_q^2)$$

$$\text{Cov}(Y_t, Y_{t-1}) = \sigma^2[\theta_1 + \theta_1\theta_2 + \theta_2\theta_3 + \dots + \theta_{q-1}\theta_q]$$

$$\text{Cov}(Y_t, Y_{t-2}) = \sigma^2[\theta_2 + \theta_1\theta_3 + \theta_2\theta_4 + \dots + \theta_{q-2}\theta_q]$$

\vdots

$$\text{Cov}(Y_t, Y_{t-q}) = \sigma^2[\theta_q]$$

$$\text{Cov}(Y_t, Y_{t-j}) = 0 \text{ for } j > q$$

Given finite order q , pure MA(q) is always stationary.

Stationary MA(∞)

However, stationarity cannot be guaranteed for an MA process with infinite order.

Even convergence of θ_q or $\theta_q \rightarrow 0$ as $q \rightarrow \infty$

does not imply stationarity.

Stationary AR(1) ----(2)

$$\begin{aligned} E(Y_t) &= \mu + \phi_1 \mu + \phi_1^2 \mu + \dots \\ &+ \dots + \phi_1^2 E(\varepsilon_{t-2}) + \phi_1 E(\varepsilon_{t-1}) + E(\varepsilon_t) \\ &= \mu + \phi_1 \mu + \phi_1^2 \mu + \dots \\ &= \frac{\mu}{1 - \phi_1} \quad \text{if } |\phi_1| < 1 \\ &\implies \text{Mean-stationary} \end{aligned}$$

Stationary AR(1) ----(1)

Given $Y_t = \mu + \phi_1 Y_{t-1} + \varepsilon_t$

$$Y_t = \mu + \phi_1 (\mu + \phi_1 Y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t$$

$$= \mu + \phi_1 \mu + \phi_1^2 Y_{t-2} + \phi_1 \varepsilon_{t-1} + \varepsilon_t$$

$$= \mu + \phi_1 \mu + \phi_1^2 \mu + \dots$$

$$\left(\dots + \phi_1^2 \varepsilon_{t-2} + \phi_1 \varepsilon_{t-1} + \varepsilon_t \right)$$

MA(∞)

Stationary AR(1) ----(3)

$$\begin{aligned} V(Y_t) &= \dots + \phi_1^4 V(\varepsilon_{t-2}) + \phi_1^2 V(\varepsilon_{t-1}) + V(\varepsilon_t) \\ &= \dots + \phi_1^4 \sigma^2 + \phi_1^2 \sigma^2 + \sigma^2 \\ &= \frac{\sigma^2}{1 - \phi_1^2} \quad \text{if } |\phi_1| < 1 \end{aligned}$$

\implies Variance-stationary

Stationary AR(1) ----(4)

$$\begin{aligned} Cov(Y_t, Y_{t-1}) &= \dots + \phi_1^3 V(\varepsilon_{t-2}) + \phi_1 V(\varepsilon_{t-1}) \\ &= \dots + \phi_1^3 \sigma^2 + \phi_1 \sigma^2 \\ &= \frac{\sigma^2}{1 - \phi_1^2} \phi_1 \quad \text{if } |\phi_1| < 1 \end{aligned}$$

Stationary AR(1) ----(6)

With linear time trend

$$\begin{aligned} Y_t &= \mu + \delta t + \phi_1 Y_{t-1} + \varepsilon_t \\ Y_t &= \mu + \delta t + \phi_1 (\mu + \delta(t-1) + \phi_1 Y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t \\ &= \mu + \phi_1 \mu + \phi_1^2 \mu + \dots \\ &\quad - (\phi_1 + 2\phi_1^2 + 3\phi_1^3 + \dots) \delta \\ &\quad + (1 + \phi_1 + \phi_1^2 + \dots) \delta t \\ &\quad + \dots + \phi_1^2 \varepsilon_{t-2} + \phi_1 \varepsilon_{t-1} + \varepsilon_t \end{aligned}$$

←MA(∞)

Stationary AR(1) ----(5)

$$\begin{aligned} Cov(Y_t, Y_{t-j}) &= \dots + \phi_1^{j+2} V(\varepsilon_{t-j}) + \phi_1^j V(\varepsilon_{t-j}) \\ &= \dots + \phi_1^{j+2} \sigma^2 + \phi_1^j \sigma^2 \\ &= \frac{\sigma^2}{1 - \phi_1^2} \phi_1^j \quad \text{if } |\phi_1| < 1 \end{aligned}$$

==> Covariance-stationary

==> also converge to zero

Stationary AR(1) ----(7)

For $|\phi_1| < 1$

$$E(Y_t) = \frac{\mu}{1 - \phi_1} - \frac{\phi_1 \delta}{(1 - \phi_1)^2} + \frac{\delta}{1 - \phi_1} t$$

==> mean non-stationary

but still variance and

covariance-stationary

Stationary AR(1) ----(8)

Note that AR(1) with linear time trend is equivalent to

$$Y_t - \frac{\delta}{1-\phi_1}t = \mu + \delta - \frac{\delta}{1-\phi_1} + \phi_1 \left(Y_{t-1} - \frac{\delta}{1-\phi_1}(t-1) \right) + \varepsilon_t$$

Stationary AR(1) ----(10)

Note that

$$AR(1) \implies MA(\infty)$$

but $AR(1) \not\Leftarrow MA(\infty)$

Stationary AR(1) ----(9)

Can prove that the de-trended Y or

$$Z_t = Y_t - \frac{\delta}{1-\phi_1}t$$

is weakly stationary.

$\implies Y$ is defined as trend-stationary

Random Walk (1)

AR(1) with $\phi_1 = 1$

Without drift

$$Y_t = Y_{t-1} + \varepsilon_t \quad \text{or} \quad \Delta Y_t = \varepsilon_t$$

With drift (intercept or constant term)

$$\Delta Y_t = \mu + \varepsilon_t$$

where $\Delta Y_t = Y_t - Y_{t-1}$

or the first difference of Y_t

Random Walk (2)

AR(1) with $\phi_1 = 1$

Note that a random walk process is non-stationary but its first difference is.

Y is also referred to as first-difference stationary.

Stationary AR(p) ----(1)

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + \varepsilon_t$$

$$Y_t = \phi_1 L Y_t + \phi_2 L^2 Y_t + \dots + \phi_p L^p Y_t + \varepsilon_t$$

$$(1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p) Y_t = \varepsilon_t$$

Stationary ARMA(1,q)

$$Y_t = \phi_1 Y_{t-1} + u_t$$

where u_t follows a stationary MA process with auto-covariance

$$\psi_j = Cov(u_t, u_{t-j})$$

Given $|\phi_1| < 1$, Y is stationary if ψ_j converges to zero

Stationary AR(p) ----(2)

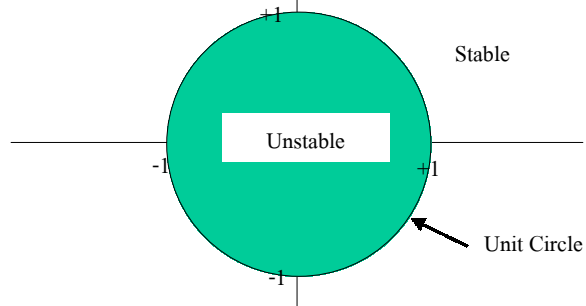
Polynomial Characteristic function of AR(p)

$$1 - \phi_1 \lambda - \phi_2 \lambda^2 - \dots - \phi_p \lambda^p = 0$$

Stability Condition: Y is stable if all the p roots of the above polynomial function (or AR roots) must lie outside the unit circle. Why? If not, Y will diverge away from its mean

Stationary AR(p) ----(3)

Note that a root could be real, complex and it could be unique or repeated.



Stationary AR(p) ----(5)

Alternative Polynomial Characteristic function of AR(p)

$$\lambda^p - \phi_1 \lambda^{p-1} - \phi_2 \lambda^{p-2} - \dots - \phi_p = 0$$

Stability Condition : all the p roots of the above polynomial function (or inverted AR roots) must lie inside the unit circle

Stationary AR(p) ----(4)

For $p=1$, $1 - \phi_1 \lambda = 0$

The only root is $\lambda = \frac{1}{\phi_1}$

Stability Condition requires that

$$|\lambda| > 1 \implies |\phi_1| < 1$$

For AR(1), Stability cond. = Stationarity

Stationary AR(p) ----(6)

$$(1 - \lambda_1 L)(1 - \lambda_2 L) \dots (1 - \lambda_p L) Y_t = \varepsilon_t$$

Given that λ_1 is inside the unit circle,

$$(1 - \lambda_2 L) \dots (1 - \lambda_p L) Y_t = (1 - \lambda_1 L)^{-1} \varepsilon_t$$

is stationary.

Given that all the p AR roots are inside the unit circle, Y is stationary

AR(1)

Stationary ARMA(p,q)

Stability condition is only a necessary condition for stationarity condition.

Need additional conditions about MA part

Stationary ARMA(p,q)

Polynomial Characteristic function for AR part

$$1 - \phi_1\lambda - \phi_2\lambda^2 - \dots - \phi_p\lambda^p = 0$$

Polynomial Characteristic function for MA part

$$1 + \theta_1\lambda + \theta_2\lambda^2 + \dots + \theta_q\lambda^q = 0$$

Stationary ARMA(p,q)

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \dots + \theta_q \varepsilon_{t-q}$$

$$Y_t = \phi_1 L Y_t + \phi_2 L^2 Y_t + \dots + \phi_p L^p Y_t + \varepsilon_t + \theta_1 L \varepsilon_t + \theta_2 L^2 \varepsilon_t + \dots + \theta_q L^q \varepsilon_t$$

$$(1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p) Y_t =$$

$$(1 + \theta_1 L + \theta_2 L^2 + \dots + \theta_q L^q) \varepsilon_t$$

Stationary ARMA(p,q)

Invertibility Conditions :

==> all the q roots of the MA polynomial function must lie outside the unit circle

Stability of AR + Invertibility of MA

====>Stationarity of ARMA

Stationary ARMA(p,q)

Alternative Polynomial Characteristic function
of MA part

$$\lambda^q - \theta_1 \lambda^{q-1} - \theta_2 \lambda^{q-2} - \dots - \theta_q = 0$$

Invertibility Condition : all the q roots of
the above polynomial function (or inverted
MA roots) must lie inside the unit circle

Box-Jenkins Method (2)

Steps

- de-trending
- identification
- estimation
- diagnosis checking

Box-Jenkins Method (1)

Major Assumptions

- Y has no trend
- Y is generated by an
ARIMA process with
unknown (p,d,q)

Box-Jenkins Method (3)

De-trending

- run OLS of Y_t against t

$$Y_t = \alpha + \beta t + \varepsilon_t$$

- use $Y_t^D = Y_t - \hat{\alpha} - \hat{\beta}t$ in the next
step

Box-Jenkins Method (4)

Identification

- set $d=0$
- plot sample correlogram of de-trended $\Delta^d Y_t$
- What is correlogram?
 - Plot of ACF (AC vs time)
 - Plot of PACF (PACF vs time)

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Box-Jenkins Method (6)

Sample Auto-Covariance

$$\hat{\gamma}_s = \frac{1}{n} \sum_{t=1}^{n-s} (Y_t^D Y_{t-s}^D), s = 0, 1, 2, \dots$$

which is estimator of γ_s .

Sample Auto-Correlation

$$\hat{\rho}_s = \frac{\hat{\gamma}_s}{\hat{\gamma}_0}, s = 1, 2, \dots$$

which is estimator of ρ_s .

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Box-Jenkins Method (5)

Population Auto-Correlation

$$\rho_s = \frac{\gamma_s}{\gamma_0}, s = 1, 2, \dots$$

$$\rho_0 = 1$$

which is population correlation between Y_t and Y_{t-s} .

All ρ_s 's are unknown parameters.

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Box-Jenkins Method (7)

Population Partial Auto-Correlation

ρ_{ss} = correlation between Y_t and Y_{t-s} , ignoring the effects of Y 's in between (if any)

Note that $\rho_{11} = \rho_1$

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Box-Jenkins Method (8)

Formula for ρ_{ss}

$$\rho_{ss} = \frac{\rho_s - \sum_{j=1}^{s-1} \rho_{s-1,j} \rho_{s-j}}{1 - \sum_{j=1}^{s-1} \rho_{s-1,j} \rho_j}, s = 1, 2, \dots$$

$$\rho_{sj} = \rho_{s-1,j} - \rho_{ss} \rho_{s-1,s-j}, j = 1, \dots, s-1$$

Need proof?

Box-Jenkins Method (10)

Rule

Y is stationary if both ACF and PACF converge to zero.

Sample Analogy

Y is believed to be stationary if both ACF and PACF converge to “zero” region.

Box-Jenkins Method (9)

Sample Analogy

$$\hat{\rho}_{11} = \hat{\rho}_1$$

$$\hat{\rho}_{22} = \frac{\hat{\rho}_2 - \hat{\rho}_{11} \hat{\rho}_1}{1 - \hat{\rho}_{11} \hat{\rho}_1}$$

$$\hat{\rho}_{21} = \hat{\rho}_{11} - \rho_{22} \rho_{11}$$

$$\hat{\rho}_{33} = \frac{\hat{\rho}_3 - \hat{\rho}_{21} \hat{\rho}_2 - \hat{\rho}_{22} \hat{\rho}_1}{1 - \hat{\rho}_{21} \hat{\rho}_1 - \hat{\rho}_{22} \hat{\rho}_2}$$

$$\hat{\rho}_{32} = \hat{\rho}_{22} - \hat{\rho}_{33} \hat{\rho}_{21}$$

$$\hat{\rho}_{31} = \hat{\rho}_{21} - \hat{\rho}_{33} \hat{\rho}_{22}$$

Box-Jenkins Method (10)

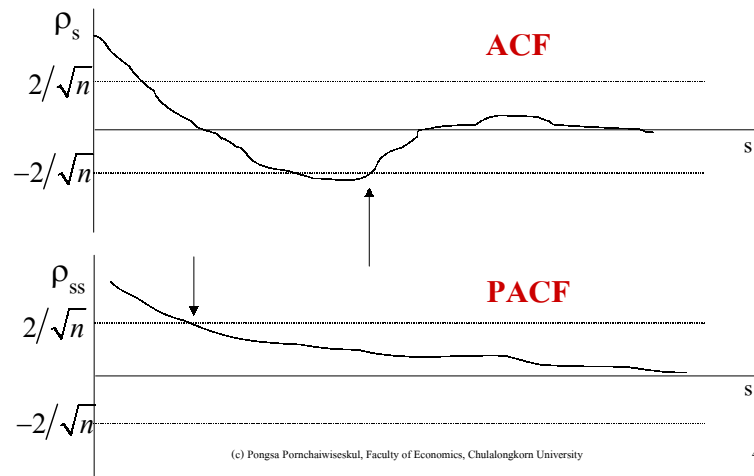
Zero Region $\hat{\rho}_s$ and $\hat{\rho}_{ss} \in 0 \pm \frac{2}{\sqrt{n}}$

$$\text{With } \rho_s = 0, \hat{\rho}_s \sim N\left(0, \frac{1 - (\hat{\rho}_s)^2}{n - 2}\right)$$

Given $\hat{\rho}_s \approx 0$ and n is large,

$$95\% \text{ CI for } \rho_s = 0 \pm 1.96 \frac{1}{\sqrt{n}}$$

Box-Jenkins Method (12)



Box-Jenkins Method (13)

Identification Rule

If $Y \sim \text{MA}(q)$, #non-zero ACF identifies q

If $Y \sim \text{AR}(p)$, #non-zero PACF identifies p

If $Y \sim \text{ARMA}(p,q)$, ACF and PACF may over-identify (p,q) .

Experiment with neighborhood of identified value of (p,q)

Box-Jenkins Method (13)

Rule

If Y is stationary, try to identify (p,q) .

Otherwise, increase order d ($d=d+1$) or try $\log(Y)$ if not conflict with underlying theories. Take difference of Y and repeat correlogram.

Box-Jenkins Method (14)

Estimation ARMA

AR part

- Cochrane-Orcutt Iterative Methods
- Durbin two-step

MA part

- FGLS

Straightforward for EViews. Use LS with lagged Y and MA terms, e.g.,

$$ls Y Y(-1) Y(-2) MA(1) MA(2)$$

Box-Jenkins Method (15)

Diagnosis Checking

- check AR roots for stability
- check MA roots for invertibility
- test for Zero Correlation of residuals (make sure that ε_t is White noise)

EViews' features and tricks

- Inverted MA roots
- Inverted AR roots (if used AR terms)
- Box-Pierce and Ljung-Box Q-tests

Box-Jenkins Method (17)

Ljung-Box's modified Q-test

$$H_0 : \rho_1 = \rho_2 = \dots = \rho_p$$

$$H_1 : \rho_1 \neq \rho_2 \neq \dots \neq \rho_p$$

$$Q = n(n+2) \sum_{s=1}^p \frac{\hat{\rho}_s^2}{n-s} \sim \chi^2(p)$$

Ljung-Box's Q is valid for small n. It has replaced Box-Pierce Q-test

Box-Jenkins Method (16)

Box-Pierce's Q-test

$$H_0 : \rho_1 = \rho_2 = \dots = \rho_p$$

$$H_1 : \rho_1 \neq \rho_2 \neq \dots \neq \rho_p$$

$$Q = n \sum_{s=1}^p \hat{\rho}_s^2 \sim \chi^2(p)$$

Proof $\sqrt{n}\hat{\rho}_s \sim N(0,1)$ for all s

$$\implies Q = n\hat{\rho}_1^2 + n\hat{\rho}_2^2 + \dots + n\hat{\rho}_p^2 \sim \chi^2(p)$$

Note that Box-Pierce's Q is not OK for small n