# 13. Recursive Identification Methods

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- Recursive prediction error method

### Introduction

Features of recursive (online) identification

- ullet  $\hat{ heta}(t)$  is computed by some 'simple modification' of  $\hat{ heta}(t-1)$
- used in central part of adaptive systems
- not all data are stored, so a small requirement on memory
- easily modified into real-time algorithms
- used in fault detection, to find out if the system has changed significantly

How to estimate time-varying parameters

- update the model regularly
- make use of previous calculations in an efficient manner
- the basic procedure is to modify the corresponding off-line method

### Desirable properties of recursive algorithms

- fast convergence
- consistent estimates (time-invariant case)
- good tracking (time-varying case)
- computationally simple

#### **Trade-offs**

- convergence vs tracking
- computational complexity vs accuracy

# Recursive least-squares method (RLS)

Recursive estimation of a constant: Consider the model

$$y(t) = b + \nu(t), \quad \nu(t)$$
 is a disturbance of variance  $\lambda^2$ 

The least-squares estimate of b is the arithmetic mean:

$$\hat{\theta}(t) = \frac{1}{t} \sum_{k=1}^{t} y(k)$$

This expression can be reformulated as

$$\hat{\theta}(t) = \hat{\theta}(t-1) + \frac{1}{t}[y(t) - \hat{\theta}(t-1)]$$

- the current estimate is equal to the previous estimate plus a correction
- the correction term is the deviation of the predicted value from what is actually observed

### RLS algorithm for a general linear model

$$y(t) = H(t)\theta + \nu(t)$$

The recursive least-squares algorithm is given by

$$\begin{array}{lcl} \hat{\theta}(t) & = & \hat{\theta}(t-1) + K(t)e(t) \\ K(t) & = & P(t)H(t)^* = & P(t-1)H(t)^*[I+H(t)P(t-1)H(t)^*]^{-1} \\ e(t) & = & y(t) - H(t)\hat{\theta}(t-1) \\ P(t) & = & P(t-1) - P(t-1)H^*(t)[I+H(t)P(t-1)H^*(t)]^{-1}H(t)P(t-1) \end{array}$$

- ullet interprete e(t) as a prediction error and K(t) as a gain factor
- ullet the update rule in P(t) has an efficient matrix inversion for scalar case

Proof of the update formula The least-square estimate is given by

$$\hat{\theta}(t) = \left(\sum_{k=1}^{t} H(t)^* H(t)\right)^{-1} \left(\sum_{k=1}^{t} H(t)^* y(t)\right)$$

Denote P(t) as

$$P(t) = \left(\sum_{k=1}^{t} H(t)^* H(t)\right)^{-1} \implies P^{-1}(t) = P^{-1}(t-1) + H(t)^* H(t)$$

Then it follows that

$$\hat{\theta}(t) = P(t) \left[ \sum_{k=1}^{t-1} H(k)^* y(k) + H(t)^* y(t) \right]$$

$$= P(t) \left[ P^{-1}(t-1)\hat{\theta}(t-1) + H(t)^* y(t) \right]$$

$$\hat{\theta}(t) = P(t) \left[ (P^{-1}(t) - H(t)^* H(t)) \hat{\theta}(t-1) + H(t)^* y(t) \right]$$

$$= \hat{\theta}(t-1) + P(t)H(t)^* \left[ y(t) - H(t) \hat{\theta}(t-1) \right]$$

To obtain the update rule for P(t), we apply the matrix inversion lemma:

$$(A + BCD)^{-1} = A^{-1} - A^{-1}B(C^{-1} + DA^{-1}B)^{-1}DA^{-1}$$

to

$$P^{-1}(t) = P^{-1}(t-1) + H(t)^*H(t)$$

where we use

$$A = P^{-1}(t-1), \quad B = H(t)^*, \quad C = I \quad D = H(t)$$

### **Initial conditions**

- $\hat{\theta}(0)$  is the initial parameter estimate
- ullet P(0) is an estimate of the covariance matrix of the initial parameter
- ullet if P(0) is small then K(t) will be small and  $\hat{ heta}(t)$  will not change much
- if P(0) is large,  $\hat{\theta}(t)$  will quickly jump away from  $\hat{\theta}(0)$
- it is common in practice to choose

$$\hat{\theta}(0) = 0, \quad P(0) = \rho I$$

where  $\rho$  is a constant

ullet using a large ho is good if the initial estimate  $\hat{ heta}(0)$  is uncertain

### **Effect of the initial values**

We simulate the following system

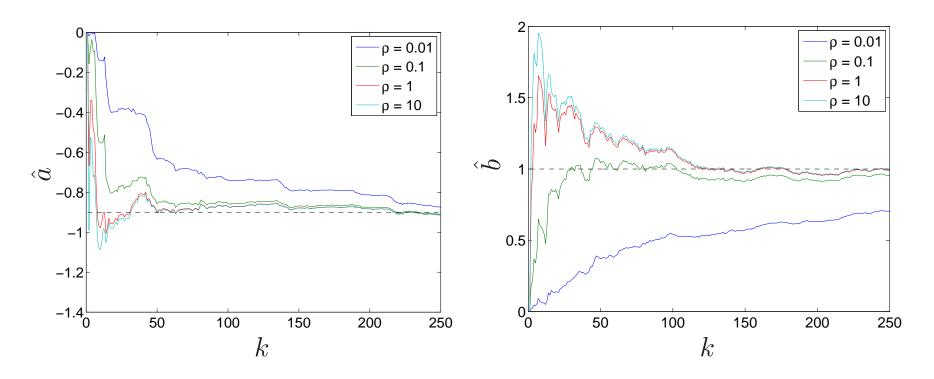
$$y(t) - 0.9y(t - 1) = 1.0u(t - 1) + \nu(t)$$

- u(t) is binary white noise
- ullet u(t) is white noise of zero mean and variance 1
- Identify the system using RLS with 250 points of data
- The parameters are initialized by

$$\hat{\theta}(0) = 0, \quad P(0) = \rho \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

for 
$$\rho = 0.01, 0.1, 1, 10$$

The graphs show the influence of the initial values



- $\bullet$  Large and moderate values of  $\rho$  (i.e.,  $\rho=1$  and  $\rho=10)$  lead to similar results
- $\bullet$  For large  $\rho$  , little confidence is given to  $\hat{\theta}(0)$  , so quick transient response
- ullet A small value of ho leads to a small K(t), so it gives a slower convergence

## **Forgetting factor**

The loss function function in the least-squares method is modified as

$$f(\theta) = \sum_{k=1}^{t} \lambda^{t-k} ||y(k) - H(k)\theta||_{2}^{2}$$

- $\bullet$   $\lambda$  is called the forgetting factor and take values in (0,1)
- the smaller the value of  $\lambda$ , the quicker the previous info will be forgotten
- the parameters are adapted to describe the newest data

The RLS method with a forgetting factor is

$$\hat{\theta}(t) = \hat{\theta}(t-1) + K(t)[y(t) - H(t)\hat{\theta}(t-1)]$$

$$K(t) = P(t)H(t)^* = P(t-1)H(t)^*[\lambda I + H(t)P(t-1)H(t)^*]^{-1}$$

$$P(t) = \frac{1}{\lambda} \left\{ P(t-1) - P(t-1)H^*(t)[\lambda I + H(t)P(t-1)H^*(t)]^{-1}H(t)P(t-1) \right\}$$

The solution  $\hat{\theta}(t)$  that minimizes  $f(\theta)$  is given by

$$\hat{\theta}(t) = \left(\sum_{k=1}^{t} \lambda^{t-k} H(k)^* H(k)\right)^{-1} \left(\sum_{k=1}^{t} \lambda^{t-k} H(k)^* y(k)\right)$$

The update formula follow analogously to RLS by introducing

$$P(t) = \left(\sum_{k=1}^{t} \lambda^{t-k} H(k)^* H(k)\right)^{-1}$$

The choice of  $\lambda$  is a trade-off between convergence and tracking performance

- $\bullet$   $\lambda$  small  $\Longrightarrow$  old data is forgotten fast, hence good tracking
- $\bullet$   $\lambda$  close to  $1 \Longrightarrow \mathsf{good}$  convergence and small variances of the estimates

## **Effect of the forgetting factor**

Consider the problem of tracking a time-varying system

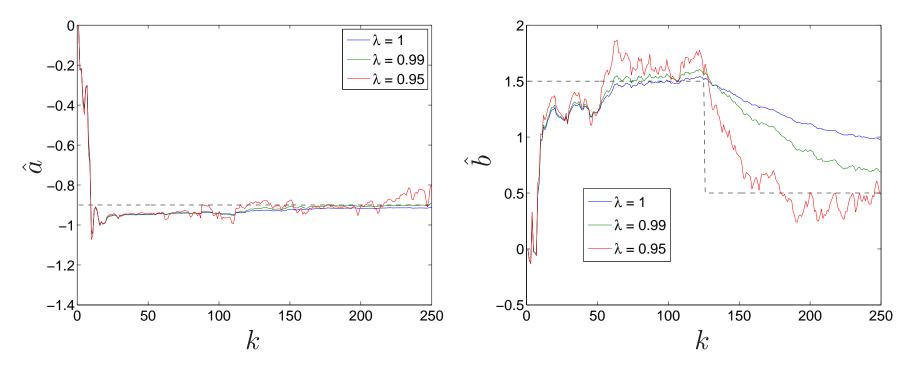
$$y(t) - 0.9y(t - 1) = b_0 u(t) + \nu(t), \quad b_0 = \begin{cases} 1.5 & t \le N/2 \\ 0.5 & t > N/2 \end{cases}$$

- u(t) is binary white noise
- ullet u(t) is white noise of zero mean and variance 1
- ullet Identify the system using RLS with 250 points of data
- The parameters are initialized by

$$\hat{\theta}(0) = 0, \quad P(0) = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$

• The forgetting factors are varied by the values  $\lambda = 1, 0.99, 0.95$ 

The graphs show the influence of the forgetting factors



A decrease in the forgetting factor leads to two effects:

- the estimates approach the true value mroe rapidly
- the algorithm becomes more sensitive to noise

As  $\lambda$  decreases, the oscillations become larger

#### In summary

- ullet one must have  $\lambda=1$  to get convergence
- ullet if  $\lambda < 1$  the parameter estimate can change quickly, and the algorithm becomes more sensitive to noise

For this reason, it is often to allow the forgetting factor to vary with time

A typically choice is to let  $\lambda(t)$  tend exponentially to 1

$$\lambda(t) = 1 - \lambda_0^t (1 - \lambda(0))$$

which can be easily implemented via a recursion

$$\lambda(t) = \lambda_0 \lambda(t-1) + (1-\lambda_0)$$

Typical values for  $\lambda_0 = 0.99$  and  $\lambda(0) = 0.95$ 

## Kalman Filter interpretation

The Kalman filter for estimating the state of a time-varying system

$$x(t+1) = A(t)x(t) + Bu(t) + \nu(t)$$
$$y(t) = C(t)x(t) + \eta(t)$$

where  $\nu(t), \eta(t)$  are independent white noise with covariances  $R_1, R_2$ , resp., is given by

$$\hat{x}(t+1) = A(t)\hat{x}(t) + B(t)u(t) + K(t)[y(t) - C(t)\hat{x}(t)]$$

$$K(t) = A(t)P(t)C^*(t)[C(t)P(t)C^*(t) + R_2]^{-1}$$

$$P(t+1) = A(t)P(t)A^*(t) + R_1$$

$$-A(t)P(t)C(t)^*[C(t)P(t)C^*(t) + R_2]^{-1}C(t)P(t)A(t)^*$$

The linear regression model

$$y(t) = H(t)\theta + \nu(t)$$

can be written as a state-space equation

$$\theta(t+1) = \theta(t) \quad (=\theta)$$
$$y(t) = H(t)\theta(t) + \nu(t)$$

Apply the Kalman filter to the state-space equation with

$$A(t) = I$$
,  $B(t) = 0$ ,  $C(t) = H(t)$ ,  $R_1 = 0$ 

When  $R_2 = I$ , it will give precisely the basic RLS algorithm in page 13-5

The tracking capability is affected by  $R_2$ 

### Recursive instrument variable method

The IV estimate of a scalar linear system

$$y(t) = H(t)\theta + \nu(t)$$

is given by

$$\hat{\theta}(t) = \left[\sum_{k=1}^{t} Z(k)^* H(k)\right]^{-1} \left[\sum_{k=1}^{t} Z(k)^* y(k)\right]$$

The IV estimate can be computed recursively as

$$\hat{\theta}(t) = \hat{\theta}(t-1) + K(t)[y(t) - H(t)\hat{\theta}(t-1)] 
K(t) = P(t)Z(t)^* = P(t-1)Z(t)^*[I + H(t)P(t-1)Z(t)^*] 
P(t) = P(t-1) - P(t-1)Z(t)^*[I + H(t)P(t-1)Z(t)^*]^{-1}H(t)P(t-1)$$

(analogous proof to RLS by using  $P(t) = (\sum_{k=1}^{t} Z(k)^*H(k))^{-1}$ )

## Recursive prediction error method

We will use the cost function

$$f(t,\theta) = \frac{1}{2} \sum_{k=1}^{t} \lambda^{t-k} e^*(k,\theta) W e(k,\theta)$$

where  $W \succ 0$  is a weighting matrix

- For  $\lambda=1$ ,  $f(\theta)=\mathbf{tr}(WR(\theta))$  where  $R(\theta)=\frac{1}{2}\sum_{k=1}^t e(k,\theta)e^*(k,\theta)$
- ullet The off-line estimate of  $\hat{ heta}$  cannot be found analytically (except for the LS case)
- It is not possible to derive an exact recursive algorithm
- Some approximation must be used, and they hold exactly for the LS case

Main idea: Assume that

- $\hat{\theta}(t-1)$  minimizes  $f(t-1,\theta)$
- the minimum point of  $f(t,\theta)$  is close to  $\hat{\theta}(t-1)$

Using a second-order Taylor series approximation around  $\hat{\theta}(t-1)$  gives

$$f(t,\theta) \approx f(t,\hat{\theta}(t-1)) + \nabla f(t,\hat{\theta}(t-1))^* (\theta - \hat{\theta}(t-1)) + \frac{1}{2} [\theta - \hat{\theta}(t-1)]^* \nabla^2 f(t,\hat{\theta}(t-1)) [\theta - \hat{\theta}(t-1)]$$

Minimize this w.r.t.  $\theta$  and let the minimizer be  $\hat{\theta}(t)$ :

$$\hat{\theta}(t) = \hat{\theta}(t-1) - [\nabla^2 f(t, \hat{\theta}(t-1))]^{-1} \nabla f(t, \hat{\theta}(t-1))$$

(Newton-Raphson step)

We must find  $\nabla f(t, \hat{\theta}(t-1))$  and  $P(t) = [\nabla^2 f(t, \hat{\theta}(t-1))]^{-1}$ 

**Details:** To proceed, the gradient of  $f(t,\theta)$  w.r.t  $\theta$  are needed

$$f(t,\theta) = \lambda f(t-1,\theta) + \frac{1}{2}e(t,\theta)^* W e(t,\theta)$$

$$\nabla f(t,\theta) = \lambda \nabla f(t-1,\theta) + e(t,\theta)^* W \nabla e(t,\theta)$$

$$\nabla^2 f(t,\theta) = \lambda \nabla^2 f(t-1,\theta) + \nabla e(t,\theta)^* W \nabla e(t,\theta) + e(t,\theta)^* W \nabla^2 e(t,\theta)$$

### First approximations:

• 
$$\nabla^2 f(t-1,\hat{\theta}(t-1)) = \nabla^2 f(t-1,\hat{\theta}(t-2))$$
 ( $\nabla^2 f$  varies slowly with  $\theta$ )

•  $e(t,\theta)^*W\nabla^2e(t,\theta)$  is negligible

After inserting the above equations to

$$\hat{\theta}(t) = \hat{\theta}(t-1) - [\nabla^2 f(t, \hat{\theta}(t-1))]^{-1} \nabla f(t, \hat{\theta}(t-1))$$

we will have

$$\hat{\theta}(t) = \hat{\theta}(t-1) - [\nabla^2 f(t, \hat{\theta}(t-1))]^{-1} [e(t, \hat{\theta}(t-1))^* W \nabla e(t, \hat{\theta})(t-1)]$$

$$\nabla^2 f(t, \hat{\theta}(t-1)) = \lambda \nabla^2 f(t-1, \hat{\theta}(t-2)) + \nabla e(t, \hat{\theta}(t-1))^* W \nabla e(t, \hat{\theta}(t-1))$$

(still not suited well as an online algorithm due to the term  $e(t, \hat{\theta}(t-1))$ 

### **Second approximations:** Let

$$e(t) \approx e(t, \hat{\theta}(t-1)), \quad H(t) \approx -\nabla e(t, \hat{\theta}(t-1))$$

(the actual way of computing these depends on model structures), then

$$\hat{\theta}(t) = \hat{\theta}(t-1) + P(t)H^*(t)We(t)$$

where we denote  $P(t) = [\nabla^2 f(t, \hat{\theta}(t-1))]^{-1}$  which satisfies

$$P^{-1}(t) = \lambda P^{-1}(t-1) + H(t)^* W H(t)$$

Apply the matrix inversion lemma to the recursive formula of  $P^{-1}(t)$ 

we arrive at recursive prediction error method (RPEM)

### **Algorithm:**

$$\begin{split} \hat{\theta}(t) &= \hat{\theta}(t-1) + K(t)e(t) \\ K(t) &= P(t)H(t)^* \\ P(t) &= \frac{1}{\lambda} \left\{ P(t-1) - P(t-1)H(t)^* [\lambda W^{-1} + H(t)P(t-1)H(t)^*]^{-1} P(t-1) \right\} \end{split}$$

where the approximations

$$e(t) \approx e(t, \hat{\theta}(t-1)), \quad H(t) \approx -\nabla e(t, \hat{\theta}(t-1))$$

depend on the model structure

## **Example of RPEM: ARMAX models**

Consider the scalar ARMAX model

$$A(q^{-1})y(t) = B(q^{-1})u(t) + C(q^{-1})\nu(t)$$

where all the polynomials have the same order

$$A(q^{-1}) = 1 + a_1 q^{-1} + \dots + a_n q^{-n}$$

$$B(q^{-1}) = b_1 q^{-1} + \dots + b_n q^{-n}$$

$$C(q^{-1}) = 1 + c_1 q^{-1} + \dots + c_n q^{-n}$$

Define

$$\tilde{y}(t,\theta) = \frac{1}{C(q^{-1})}y(t), \quad \tilde{u}(t,\theta) = \frac{1}{C(q^{-1})}u(t), \quad \tilde{e}(t,\theta) = \frac{1}{C(q^{-1})}e(t)$$

We can derive the following relations

$$e(t,\theta) = \frac{A(q^{-1})y(t) - B(q^{-1})u(t)}{C(q^{-1})}$$

$$\nabla e(t,\theta) = (\tilde{y}(t-1,\theta), \dots, \tilde{y}(t-n,\theta), -\tilde{u}(t-1,\theta), \dots, -\tilde{u}(t-n,\theta), -\tilde{u}(t-1,\theta), \dots, -\tilde{u}(t-n,\theta), -\tilde{u}(t-1,\theta), \dots, -\tilde{u}(t-n,\theta), -\tilde{u}(t-n,\theta))$$

To compute  $e(t,\theta)$ , we need to process all data up to time t

We use the following approximations

$$e(t,\theta) \approx e(t) = y(t) + \hat{a}_1(t-1)y(t-1) + \dots + \hat{a}_n(t-1)y(t-n)$$
$$-\hat{b}_1(t-1)u(t-1) - \dots - \hat{b}_n(t-1)u(t-n)$$
$$-\hat{c}_1(t-1)e(t-1) - \dots - \hat{c}_n(t-1)e(t-n)$$

$$-\nabla e(t,\theta) \approx H(t) = (-\bar{y}(t-1), \dots, -\bar{y}(t-n), \\ \bar{u}(t-1), \dots, \bar{u}(t-n), \bar{e}(t-1), \dots, \bar{e}(t-n))$$

where

$$\bar{y}(t) = y(t) - \hat{c}_1(t)\bar{y}(t-1) - \dots - \hat{c}_n(t)\bar{y}(t-n)$$

$$\bar{u}(t) = u(t) - \hat{c}_1(t)\bar{u}(t-1) - \dots - \hat{c}_n(t)\bar{u}(t-n)$$

$$\bar{e}(t) = e(t) - \hat{c}_1(t)\bar{e}(t-1) - \dots - \hat{c}_n(t)\bar{e}(t-n)$$

## Comparison of recursive algorithms

We simulate the following system

$$y(t) = \frac{1.0q^{-1}}{1 - 0.9q^{-1}}u(t) + \nu(t)$$

- $u(t), \nu(t)$  are indepentdent white noise with zero mean and variance 1
- we use RLS,RIV, RPEM to identify the system

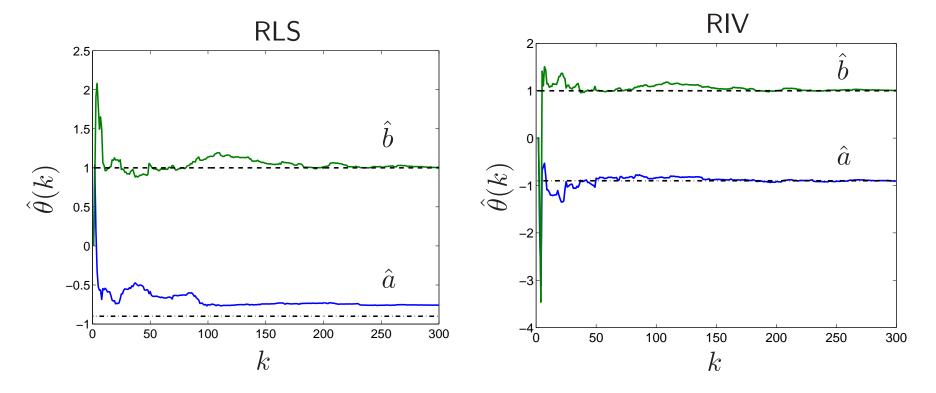
#### The model structure for RLS and RIV:

$$y(t) + ay(t-1) = bu(t-1) + \nu(t), \quad \theta = (a, b)$$

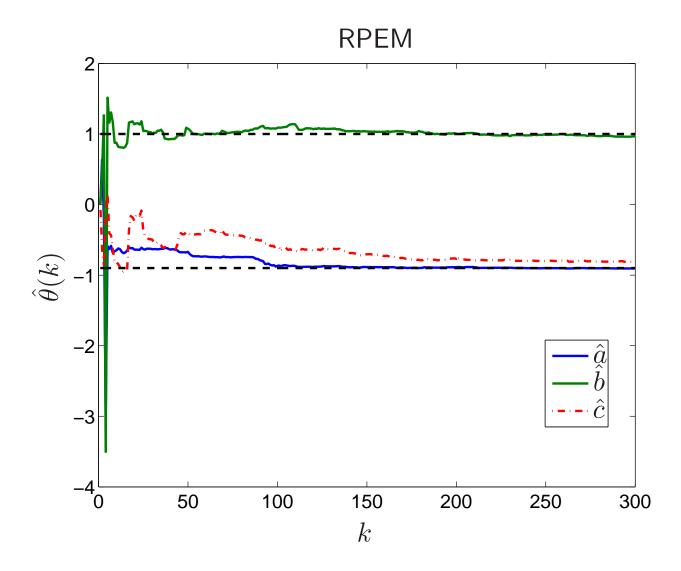
#### The model structure for RPEM:

$$y(t) + ay(t-1) = bu(t-1) + \nu(t) + c\nu(t-1), \quad \theta = (a, b, c)$$

#### **Numerical results**



- RLS does not give consistent estimates for systems with correlated noise
- This is because RLS is equivalent to an off-line LS algorithm
- In contrast to RLS, RIV gives consistent estimates
- This result follows from that RIV is equivalent to an off-line IV method



- $\bullet$  RPEM gives consistent estimates of a,b,c
- $\bullet$  The estimates  $\hat{a}$  and  $\hat{b}$  converge more quickly and  $\hat{c}$

# Common problems for recursive identification

- Excitation
- Estimator windup
- P(t) becomes indefinite

**Excitation** it is important that the input is persistently excitation of sufficiently high order

## **Estimator windup**

Some periods of an identification experiment exhibit poor excitation Consider when H(t)=0 in the RLS algorithm, then

$$\hat{\theta}(t) = \hat{\theta}(t-1), \quad P(t) = \frac{1}{\lambda}P(t-1)$$

- $\hat{\theta}$  becomes constant as t increases
- $\bullet$   $\,P$  increases exponentially with time for  $\lambda < 1$
- When the system is excited again  $(H(t) \neq 0)$ , the gain

$$K(t) = P(t)H(t)^*$$

will be very large and causes an abrupt change in  $\hat{ heta}$ 

• This is referred to as *estimator windup* 

**Solution:** Do not update P(t) if we have poor excitation

# P(t) indefinite

P(t) represents a covariance matrix

Therefore, it must be symmetric and positive definite

Rounding error may accumulate to make P(t) indefinite

This will make the estimate diverge

The solution is to note that every positive definite matrix can be factorized as

$$P(t) = S(t)S(t)^*$$

and rewrite the algorithm to update S(t) instead

### References

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